Effects of frequency and similarity neighborhoods on pharmacists' visual perception of drug names

Bruce L. Lambert[a,b,*], Ken-Yu Chang[c], Prahlad Gupta[d]

[a] Department of Pharmacy Administration, University of Illinois at Chicago, Chicago, IL, USA
[b] Department of Pharmacy Practice, University of Illinois at Chicago, Chicago, IL, USA
[c] Center for Pharmaceutical Appraisal and Outcomes Research, Abbott Laboratories, USA
[d] Department of Psychology, University of Iowa, USA

Abstract

To minimize drug name confusion errors, regulators, drug companies, and clinicians need tools that help them predict which names are most likely to be involved in confusions. Two experiments, carried out in the United States, examined the effects of stimulus frequency (i.e., how frequently a target name is prescribed), neighborhood frequency (i.e., how frequently prescribed are the "neighbors" of the target name), and neighborhood density (how many names are within a fixed distance of the target name) on the probability of pharmacists making an error in a visual perceptual identification task. In both experiments, the task was to correctly identify a series of blurry drug names after a 3 s presentation on a computer monitor. In the first experiment, 45 pharmacists viewed 160 typewritten names, incorrectly identifying 60.6% of them. Random effects regression revealed a significant beneficial effect of stimulus frequency and a detrimental effect of neighborhood density. Significant two-way interactions were observed between stimulus frequency and neighborhood density and neighborhood frequency and neighborhood density. In the second experiment, 37 pharmacists viewed 156 handwritten drug names, incorrectly identifying 45.7%. Random effects regression revealed significant main effects of stimulus frequency and neighborhood density. These were contained within a significant three-way interaction: The interaction between stimulus frequency and neighborhood density was present at high but not low neighborhood frequency. Objectively measurable frequency and neighborhood characteristics have predictable effects on errors in pharmacists' visual perception. Organizations that coin and evaluate drug names, as well as hospitals, pharmacies, and health systems, should consider these characteristics when assessing visually confusing names.
© 2003 Elsevier Science Ltd. All rights reserved.

Keywords: USA; Drug names; Patient safety; Visual perception; Pharmacists; Medication error

Introduction

Recent estimates suggest that medical errors of all types may cause the death of between 44,000 and 98,000 hospitalized patients in the United States each year (Kohn, Corrigan, & Donaldson, 2000). Errors involving

medication cause the death of one person every day in the US, and injure more than a million more each year (US Food and Drug Administration, 2001). Confusions between drug names that look and sound alike account for between 15% and 25% of reported medication errors in the US (US Pharmacopeia, 1995, 1996, 2001). Similarity between drug names can cause errors in short-term memory as well as in visual and auditory perception (Brodell, Helms, KrishnaRao, & Bredle, 1997; Lambert, Chang, & Lin, 2001b; Lambert, Chang, Lin, & Gupta, 2000; Lambert, Lin, Gandhi, & Chang, 1999; Luce & Pisoni, 1998; Luce, 1959). This

*Corresponding author. Department of Pharmacy Administration, University of Illinois at Chicago, 833 South Wood Street, Chicago, IL 60612-7231, USA. Tel.: +1-312-996-2411; fax: +1-312-996-0868.

E-mail address: lambertb@uic.edu (B.L. Lambert).

investigation examined the effect of similarity and prescribing frequency on pharmacists' ability to accurately identify blurry, briefly presented, handwritten and typewritten drug names. Two research questions provided the motivation for the experiments that follow:

(1) To what extent is a pharmacist's ability to identify a target drug name affected by the prescribing frequency of the target, the number of names similar to the target, and the prescribing frequency of the similar names?

(2) To what extent do the characteristics in Q1 have different effects depending on whether the drug name is handwritten or typewritten?

Theoretical background

Activation-competition models of visual perception

The "interactive activation" framework has been very influential in the development of cognitive psychology theories of visual word recognition, following a seminal formulation of this approach by McClelland and Rumelhart (McClelland & Rumelhart, 1981; Rumelhart & McClelland, 1982). In such a model, words are represented in memory as networks of nodes connected by excitatory and inhibitory links. The nodes are typically arranged in a hierarchy of levels. Lower level nodes detect particular letters. High-level nodes detect particular words. When a word stimulus is presented, the flow of activation is bottom-up from the feature level to the letter level and eventually to the word level (see Fig. 1 for a simplified schematic). However, a crucial aspect of the interactivity within the system is that, as time passes, the letter-level nodes can be reinforced via excitatory links from the word-level nodes, and vice versa. Thus there is a "top-down" influence on processing, because word-level nodes contribute to the activation of letter-level nodes. Another crucial aspect of this framework is that active nodes at a particular level try to suppress the activations of other nodes at the same level, via inhibitory connections. Thus nodes at a particular level "compete" to be active (Grainger & Dijkstra, 1996; Grainger & Jacobs, 1996).

This original formulation of the interactive activation theory did not aim to capture the distributional statistics of word groups. That is, it neither precisely reflected the fact that a word like "cat" has many similar words (which will therefore compete with it for activation at the word level) whereas a word like "atrocity" has few similar words (and therefore less competition at the word level), nor did it capture the frequencies of different words. Recent work has extended the basic ideas of the interactive activation framework to creating models that can represent word frequency, and that can investigate how words in a "neighborhood" interact in complex ways, leading sometimes to inhibition of each other, and sometimes to facilitation, under certain circumstances (Andrews, 1997; Dell, Schwartz, Martin, Saffran, & Gagnon, 1997; Grainger & Jacobs, 1996; Jacobs & Grainger, 1994; Mathey, 2001; Perea & Rosa, 2000). The following two sections summarize the observed effects of frequency and neighborhood characteristics on a variety of word recognition tasks.

Frequency effects on visual perception of words

One of the oldest and most consistent findings in psycholinguistics is that word frequency (normally defined as the frequency with which a word appears in print) enhances word recognition. Nearly 50 years ago, Solomon and others showed that as word frequency
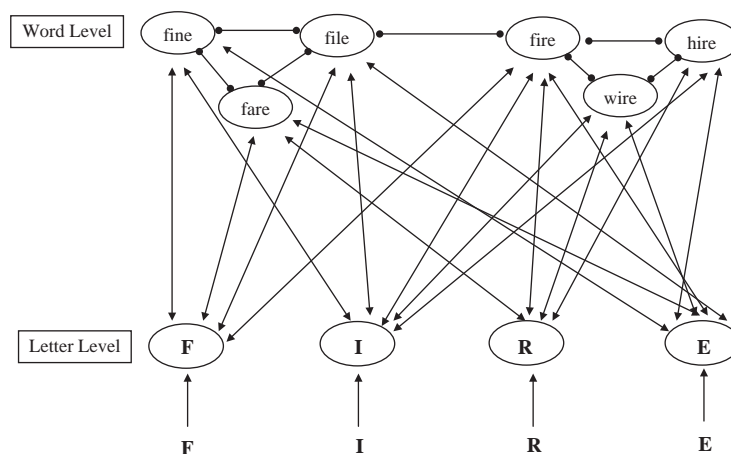

Fig. 1. Interactive activation model of visual word perception. Arrows represent excitatory links. Filled circles represent inhibitory links. See, e.g., McClelland and Rumelhart (1981).

increases, the time needed to correctly identify a briefly presented word decreases (Howes & Solomon, 1951; Solomon & Postman, 1952). Similar effects have been repeatedly demonstrated in a variety of related tasks, including lexical decision (i.e., time needed to decide whether a stimulus is a word or nonsense string), perceptual identification (i.e., identifying a briefly presented or degraded word), and naming (i.e., time needed to pronounce a visually presented word) (Balota, 1994; Grainger & Dijkstra, 1996; Grainger & Segui, 1990; Monsell, 1991; Whitlow & Cebollero, 1989). Although this study focuses on visual perception, it is worth noting that analogous frequency effects have been found for perceptual identification of auditorily presented words as well (Luce, Pisoni, & Goldinger, 1990). The precise mechanism by which frequency exerts its effects is still controversial. The main ideas are either (a) that high-frequency word nodes have higher resting activation levels and hence are quicker to exceed a recognition threshold or (b) that frequency biases decision processes independent of activation levels (Andrews, 1997; Grainger & Dijkstra, 1996; Grainger & Jacobs, 1996). Regardless of the precise mechanism, it is clear that in perceptual identification tasks, more frequent words are identified more quickly and accurately than less frequent words. This leads to our first hypothesis:

Hypothesis 1. High-frequency drug names will be identified more accurately than low-frequency drug names in a visual perceptual identification task.

Neighborhood effects on the visual perception of words

Another consistent finding is that word recognition is affected by the properties of a word's neighborhood (Andrews, 1997; Havens & Foote, 1963; Mathey, 2001; Perea & Rosa, 2000). In this context, 'neighborhood' refers to the set of words that are within some similarity boundary of the target word. The most common operational definition of orthographic (i.e., spelling) neighborhood is the set of all words of the same length that share all but one letter in the same position as the target word (Coltheart, Davelaar, Jonasson, & Besner, 1977). For example, in terms of this definition, orthographic (i.e. written) neighbors of the word 'dog' include 'cog', 'dig', and 'dot'. The set of all such words differing in only one letter from 'dog' makes up the orthographic neighborhood of dog. (As discussed below, this definition will prove to be too strict when working with drug names.) Two specific features of a word's neighborhood have received most of the experimental attention: *neighborhood frequency* and *neighborhood density* (sometimes called neighborhood size or simply N). Neighborhood frequency refers to how frequent the words in the

neighborhood are. One common operational definition is to assign neighborhood frequency the value 1 if at least one of the neighbors has a higher frequency than the word itself, and otherwise to assign neighborhood frequency a value of 0 (Andrews, 1997). Generally speaking, words with at least one higher frequency neighbor are harder to recognize than words with no higher frequency neighbors (Newbigging, 1961). In a recent review of 16 published papers that studied neighborhood effects on a variety of tasks, Andrews notes that neighborhood frequency inhibited recognition in 14 of 23 experiments (Andrews, 1997). The effects were null in 5 experiments and facilitatory in 4 experiments (Andrews, 1997). In perceptual identification, the task most relevant to the current experiments, 3 out of 3 experiments found that neighborhood frequency inhibited perceptual identification (Andrews, 1997; Carreiras, Perea, & Grainger, 1997; Grainger & Jacobs, 1996; Grainger & Segui, 1990). This leads to our second hypothesis:

Hypothesis 2. Drug names with low neighborhood frequency will be identified more accurately than names with high neighborhood frequency in a visual perceptual identification task.

The effects of neighborhood density, on the other hand, depend on the task (Andrews, 1997). For lexical decision and naming tasks, neighborhood density facilitates recognition. That is, words with many neighbors are easier to name and recognize as true words than words with few neighbors. In contrast, for the perceptual identification task that is our main focus, neighborhood density inhibits performance. Under brief or degraded stimulus conditions, performance is slower and less accurate for words with many neighbors than it is for words with few neighbors (Andrews, 1997; Carreiras et al., 1997; Snodgrass & Mintzer, 1993). Thus, our third hypothesis is as follows:

Hypothesis 3. Drug names with low neighborhood density will be identified more accurately than names with high neighborhood density in a visual perceptual identification task.

Experiment 1: Perceptual identification of typewritten drug names

Methods

Design

This experiment was designed to examine the effect of prescribing frequency, neighborhood frequency and neighborhood density on the probability of a pharmacist

making an error in a visual perceptual identification task. Participants viewed a series of noise-masked, typewritten drug names as they were briefly presented on a computer monitor. The task was to correctly identify the presented name by typing it into a provided text box. All experiments were approved in advance by the local institutional review board, and all participants orally consented to participate.

Participants

Forty-five licensed, practicing pharmacists participated in Experiment 1. Participants were recruited from among the attendees at the 2000 annual meeting of the American Pharmaceutical Association in Washington, DC. Individuals were not paid for their participation.

Stimulus materials

One hundred sixty (1 6 0) three-syllable drug names were selected to fill the cells of a $2 \times 2 \times 2$ stratified sampling design, where the strata were stimulus frequency (high/low), neighborhood frequency (high/low), and neighborhood density (high/low) (see Figs. 2 and 3 and Tables 1–3). Names and prescribing frequencies were obtained from the drug databases contained within the US National Ambulatory Medical Care Survey (NAMCS) and the US National Hospital Ambulatory Medical Care Survey (NHAMCS) for the years 1992–1996 (National Center for Health Statistics, 2001a; National Center for Health Statistics, 2001b). NAMCS and NHAMCS surveys are based on a nationally representative sample of outpatient physicians working in traditional outpatient offices (NAMCS) as well as hospital emergency departments and outpatient clinics (NHAMCS). NAMCS surveys exclude anesthesiologists, radiologists and pathologists. NHAMCS excludes
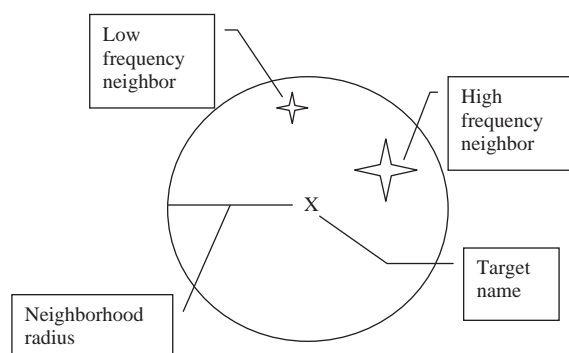

Fig. 2. Schematic diagram of a drug name's orthographic neighborhood. The size of the center X represents stimulus frequency (i.e., the prescribing frequency of the target name). The stars in the circle represent neighbors within a fixed neighborhood radius (e.g., 3 edits). The size of each star represents the prescribing frequency of that name. The ability to compete for perceptual identification increases as prescribing frequency increases. See text for details.
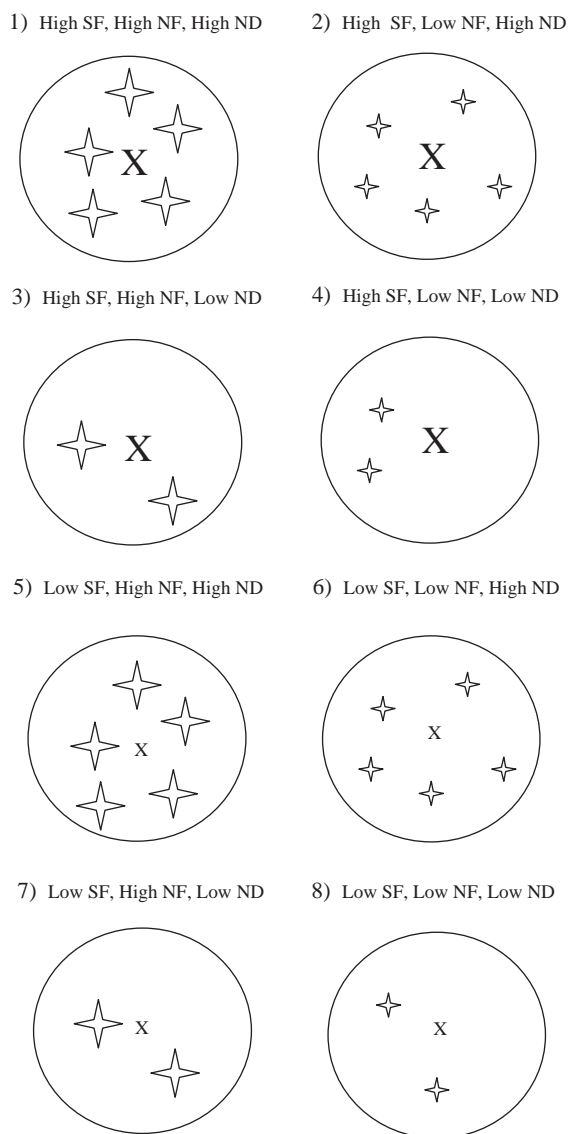

Fig. 3. Schematic diagram of 8 different combinations of stimulus frequency (SF), neighborhood frequency (NF) and neighborhood density (ND). These 8 combinations correspond to the 8 cells of the experimental design.

federal, military, and Veterans Administration Hospitals. Prescribing frequencies were based on national estimates derived by weighting raw frequencies to reflect the probabilistic sampling design of NAMCS and NHAMCS, respectively. Henceforth, when we refer to stimulus frequency, we are referring to the $\log_{10}$ of the 5-year (1992–1996) cumulative frequencies from the combined NAMCS and NHAMCS data.

Operational definitions of neighborhood, neighborhood frequency and neighborhood density. Both neighborhood frequency and neighborhood density presume some

Table 1
Stimulus names for Experiments 1 and 2

Cell	Name	SF	NF	ND
1	Trilafon	6.29	6.60	4
1	Rifampin	5.96	6.36	4
1	Depakene	5.69	6.15	4
1	Aclovate	6.15	6.12	3
1	Tobradex	6.82	6.94	3
1	Ocufen	5.85	6.49	4
1	Insulin	6.55	6.54	3
1	Betoptic	6.72	6.65	3
1	Lotensin	6.79	6.49	3
1	Zonalon	5.65	6.19	5
1	Librium	6.44	6.49	3
1	Metrogel	6.50	6.15	5
1	Adalat	6.54	6.31	5
1	Ativan	7.28	6.32	5
1	Nicotrol	5.91	6.88	3
1	calcium	6.32	6.61	5
1	Mellaril	6.66	6.64	3
1	Nolvadex	6.26	6.18	5
1	Loestrin	6.35	6.10	3
1	Anaprox	7.13	6.59	3
2	Robinul	5.81	5.22	3
2	Capozide	6.08	4.99	4
2	Topicort	6.51	4.71	3
2	Verelan	6.72	4.91	4
2	Enduron	5.69	4.13	4
2	Proventil	7.66	5.11	3
2	Imuran	6.47	5.24	3
2	Sinemet	6.63	4.67	3
2	Tetramune	6.72	4.64	3
2	Kenalog	6.87	5.16	5
2	Ventolin	7.62	5.19	4
2	Catapres	6.82	5.14	4
2	Clonidine	6.74	4.92	5
2	Pramasone	5.69	4.91	4
2	Dyazide	7.45	3.66	4
2	Dilantin	7.18	4.96	4
2	Duratuss	6.13	5.08	4
2	Tagamet	7.34	5.43	3
2	Deltasone	6.09	5.30	3
2	Serentil	5.76	4.77	4
3	Micronase	7.25	6.79	2
3	Wycillin	5.77	6.25	2
3	Prinivil	6.95	6.29	1
3	Clozaril	5.91	6.24	2
3	Stelazine	6.29	6.08	2
3	Humibid	6.67	6.15	2
3	Florinef	5.99	6.03	2
3	Pancrease	5.78	7.16	1
3	Serevent	6.55	6.90	2
3	Acular	5.78	6.26	2
3	Restoril	6.81	6.72	2
3	Diprosone	5.84	6.32	2
3	Natalins	6.44	6.72	2
3	Naprosyn	7.51	6.06	2
3	Ambenyl	6.07	6.21	2
3	Skelaxin	6.38	6.29	1
3	Tegretol	6.99	6.51	2

Table 1 (*continued*)

Cell	Name	SF	NF	ND
3	Pondimin	6.25	6.34	2
3	Isordil	7.17	6.27	2
3	Vistaril	7.02	6.51	2
4	Reserpine	5.82	3.10	1
4	Dexedrine	5.99	3.86	2
4	Phentermine	6.40	3.82	1
4	Flexeril	7.28	4.36	1
4	Unasyn	5.74	3.64	2
4	Tolinase	5.96	5.23	2
4	Cyclogyl	6.32	3.84	1
4	Prazosin	5.99	5.06	1
4	Elimite	6.01	5.28	2
4	Nicoderm	6.09	4.28	2
4	Ocuflox	6.02	5.85	1
4	Provera	7.41	5.22	2
4	Pamelor	6.87	4.55	2
4	Tenormin	7.40	3.21	1
4	Bellergal	5.71	3.01	1
4	Depakote	6.69	5.50	2
4	Fosamax	5.97	4.25	1
4	Monopril	6.70	4.79	2
4	Inapsine	5.81	4.47	1
4	Fulvicin	5.85	4.72	1
5	Alfenta	5.13	6.14	3
5	Idenal	3.34	6.65	4
5	Trobicin	3.06	6.41	3
5	Desferal	5.19	6.71	3
5	Ketalar	4.93	6.35	4
5	Travasol	4.38	6.43	3
5	Activase	5.25	6.73	4
5	Solatene	4.34	6.07	3
5	Parnatal	4.01	6.29	4
5	Norethin	4.28	6.50	5
5	Hexadrol	5.30	6.39	5
5	Dynabac	5.50	6.36	5
5	Betadine	4.77	6.11	5
5	Cytadren	4.91	6.35	3
5	Catarase	4.23	6.88	3
5	Pavulon	4.61	6.39	3
5	Altoco	4.83	6.10	4
5	Virilon	4.40	6.03	5
5	Moderil	4.19	6.90	3
5	Regroton	4.94	6.57	3
6	Hydrocort	5.16	5.13	4
6	Pentetra	4.19	4.92	3
6	Propacet	5.60	5.18	4
6	Procamide	3.08	5.05	3
6	Isolyte	4.24	4.46	3
6	Iotuss	2.98	5.32	3
6	Mantadil	4.38	4.54	4
6	Estraval	3.55	5.28	3
6	Ancobon	2.95	5.06	3
6	Isoclor	3.74	4.15	4
6	Senokot	5.22	4.81	3
6	Panmycin	3.88	5.34	3
6	Tegison	5.36	4.68	3
6	Thiola	3.97	5.24	4
6	Fentanyl	5.64	4.71	3

Table 1 (*continued*)

Cell	Name	SF	NF	ND
6	Norinyl	5.53	5.50	3
6	Loniten	5.16	5.18	4
6	Renova	4.84	4.89	4
6	Hetrazan	2.52	4.55	3
6	Carmustine	4.71	4.36	3
7	Adipost	4.45	6.06	1
7	Dantrolene	3.85	6.52	1
7	Mivacron	4.37	7.08	2
7	Inpersol	3.44	6.95	2
7	Vistazine	2.63	6.72	2
7	Dymelor	4.48	6.91	2
7	Testoderm	4.36	6.28	2
7	Triphenyl	4.44	6.58	1
7	Tramadol	4.44	7.07	2
7	Cardilate	3.01	6.84	1
7	Indomed	3.52	6.76	2
7	Keralyt	4.66	6.58	2
7	Phenelzine	3.82	6.10	2
7	Donnagel	2.88	6.57	2
7	Nebupent	4.37	6.61	1
7	Surgicel	4.81	7.05	1
7	Maxiflor	4.83	6.46	1
7	Theovent	3.44	6.90	2
7	Benzonate	4.69	6.78	2
7	Flaxedil	4.36	7.28	1
8	Orinase	5.52	5.66	2
8	Polymox	5.40	3.14	1
8	Dobutrex	4.35	2.74	1
8	Humafac	4.57	4.42	1
8	Norcuron	4.96	3.12	1
8	Brethancer	4.71	5.53	0
8	Zefazone	4.13	4.81	2
8	Natabec	3.71	5.43	2
8	Ipecac	3.48	3.74	2
8	Regitine	5.13	5.64	2
8	Hydergine	5.55	3.81	2
8	Cefizox	5.19	3.99	2
8	Gaviscon	5.28	4.07	1
8	Pravastan	5.14	4.81	1
8	Nysolone	3.57	5.94	2
8	Bronkodyl	3.78	4.96	1
8	Esidrix	5.42	5.06	2
8	Antispas	2.43	5.30	2
8	Protaphane	2.21	2.82	1
8	Theodrine	4.14	3.89	1

Note: SF—log stimulus frequency. NF—log neighborhood frequency. ND—neighborhood density.

definition of neighborhood. As noted in the introduction, an orthographic neighborhood has typically been defined as the set of words (of the same length) that differ from the target word by only one letter in one position (Andrews, 1997). An equivalent way to describe this set of words is that they have an "edit distance" of 1 from the word itself (Stephen, 1994). This definition has most commonly been applied to single syllable, three-

letter, consonant–vowel–consonant words. This definition struck us as too restrictive because (a) most drug names are multisyllabic (Lambert, Chang, & Lin, 2001a) and (b) because documented pairs of confusing names typically differ by several letters in several positions (in fact the mean edit distance for more than 1000 reported confusing pairs was 4) (Lambert et al., 1999). Since the modal drug name in the US is three syllables, we chose to use three syllable names, and we extended the definition of neighborhood to be any name that fell within an edit distance of 3 edits from the target name (i.e., one letter insertion, deletion or substitution per syllable, a logical extension of the traditional definition) (Lambert et al., 2001a, 1999). Neighborhood frequency has typically been defined as a dichotomous variable that was equal to one if the neighborhood contained at least one word whose frequency was greater than the target word (Andrews, 1997). We chose instead to define neighborhood frequency as the mean frequency of all the words in the target word's neighborhood. This definition reflected our belief that all of the names in a neighborhood potentially compete with the target name, not just the most frequent name. Finally, neighborhood density was the number of names that fell within an edit distance of 3 from the target name.

Examples. High log stimulus frequency (SF) names (i.e., those with log SF > 7) included *Ventolin*®, *Dyazide*®, and *Provera*®. Low log SF names (log SF < 3) included *Vistazine*®, *Antispas*®, and *Protaphane*®. *Flexeril*® is an example of a name from a sparse neighborhood: it had no neighbors in the combined NAMCS/NHAMCS database. An example of a name from a dense neighborhood is *Dynabac*®, whose neighbors included: *Synalar*®, *Rynatan*®, *Dynapen*®, *Dynacirc*®, and *Dynacin*®. Another name from a high-density neighborhood is *Virilon*®, whose neighbors included: *Verelan*®, *Uridon*®, *Trilion*®, *Miradon*®, and *Daricon*®. Tables 1 and 2 together describe the frequency and neighborhood characteristics of the stimulus names.

Success of stratification. In cell-to-cell comparisons, sets of names at the same level of a stratum (e.g., all names at high levels of stimulus frequency) did not differ statistically from one another on their scores for that factor. Conversely, names from cells at different levels of a sampling stratum did differ significantly from one another. All cell-to-cell comparisons were based on Tukey's Honestly Significant Difference. Table 2 gives descriptive statistics for stimulus frequency, neighborhood frequency, and neighborhood density for all 8 cells. Table 3 shows the correlation between stimulus frequency, neighborhood frequency and neighborhood density.

Stimulus degradation. The names were degraded (i.e., masked by noise) using the degradation features built into the PsyScope® experiment program (Cohen,

Table 2
Descriptive statistics for stimulus frequency, neighborhood frequency, and neighborhood density ($n = 20$ names per cell)

			High NF				Low NF			
		Variable	Mean	SD	Min	Max	Mean	SD	Min	Max
High SF	High ND	SF	6.39	0.45	5.65	7.28	6.60	0.64	5.69	7.66
		NF	6.44	0.25	6.10	6.94	4.91	0.41	3.66	5.43
		ND	3.80	0.89	3.00	5.00	3.70	0.66	3.00	5.00
	Low ND	SF	6.47	0.55	5.77	7.51	6.30	0.56	5.71	7.41
		NF	6.40	0.31	6.03	7.16	4.40	0.82	3.01	5.85
		ND	1.85	0.37	1.00	2.00	1.45	0.51	1.00	2.00
Low SF	High ND	SF	4.58	0.63	3.06	5.50	4.33	0.98	2.52	5.64
		NF	6.42	0.26	6.03	6.90	4.92	0.37	4.15	5.50
		ND	3.75	0.85	3.00	5.00	3.35	0.49	3.00	4.00
	Low ND	SF	4.04	0.67	2.63	4.83	4.43	0.99	2.21	5.55
		NF	6.70	0.32	6.06	7.28	4.44	1.01	2.74	5.94
		ND	1.60	0.50	1.00	2.00	1.45	0.60	0.00	2.00

Overall

	Variable	Mean	SD	Max	Min
	SF	5.39	1.27	7.66	2.21
	NF	5.58	1.07	7.28	2.74
	ND	2.62	1.22	5.00	0.00

Table 3
Correlations between log stimulus frequency, log neighborhood frequency, and neighborhood density for $N = 160$ drug names

	SF	NF
NF	−0.040	—
ND	0.122	0.171*

*$p < 0.05$.

MacWhinney, Flatt, & Provost, 1993). We used foreground degradation of 50% and background degradation of 10%. Foreground degradation randomly removes pixels from the foreground image. Background degradation randomly adds black pixels to the background image. Overall, the effect resembles a typewritten name received through a fax machine with a dirty print cartridge that was running out of ink. (It was not possible to capture a screen shot of the degraded stimuli as they appeared within the PsyScope program, but interested individuals can reproduce the effect precisely by using the freely available software and the specifications given above.)

Procedure

Individuals were approached in the exhibit halls and corridors of the convention center and asked: (a) if they were currently a licensed, practicing pharmacist in the US and (b) if they were interested in participating in a study of drug name confusion. Those who agreed were directed to a meeting room where the experiment was being conducted. Participants read a consent form, filled out a brief demographic questionnaire, and then seated themselves in front of a Macintosh computer with a 17-in color monitor. The experiment began by presenting the instructions on the computer screen. Participants were told that their task was to correctly identify a series of briefly presented, blurry, typewritten drug names. Each trial began when the participant pressed any key on the keyboard. A row of capital X's then appeared at the center of the screen in 36-point Times font. When the participant pressed any key, the X's disappeared and a noise-masked drug name appeared. Each name appeared for 3 s, with its first letter capitalized, in the center of the screen in 36-point Times font. After 3 s, the name was replaced by another row of capital X's for 750 ms. Finally, a text box appeared in which the participant typed the name of the drug they thought they saw. Although participants were allowed to leave a blank, they were encouraged to generate an answer (a guess) whenever possible. Participants were urged to check the spelling before proceeding. This process was repeated until each of the 160 names had been presented.

Analysis plan

The goal of our analysis was to determine the main effects of stimulus frequency, neighborhood frequency, and neighborhood density on the probability of making an error in visual perceptual identification. We were also interested in any two- and three-way interactions between the main effects. The independent variables were (a) stimulus frequency, a continuous variable representing the $\log_{10}$ of the 5-year cumulative NAMCS/HAMCS frequency for each name; (b) neighborhood frequency, a continuous variable representing the mean log frequency of all names within an edit distance of 3 edits from the stimulus name; and (c) neighborhood density, an ordinal variable representing the number of other drug names that can be found within an edit distance of 3 edits of the stimulus name. The dependent variable was probability of error, a dichotomous variable scored as 1 if the name was identified incorrectly, and 0 if the name was identified correctly. The only control variable was trial, an ordinal variable representing the sequential position of a name within the 160 trials.

Data were analyzed using MIXOR, a system for doing mixed effects, logistic regression modeling of dichotomous and ordinal data. The mixed-effects logistic regression model accommodates nesting of experimental conditions within subjects for a binary outcome and a mixture of discrete and continuous covariates that can vary either at the level of the subject or the experimental condition (Hedeker, 1999; Hedeker & Gibbons, 1994; Hedeker & Gibbons, 1996).

Our modeling strategy included multiple steps. First we centered all variables at their mean values. The next step was to identify the correct scale for each independent and control variable. We did this by separately plotting the log odds of error as a function of each independent or control variable. If these plots were linear, terms were entered as simple linear terms. If the plot revealed an obvious nonlinearity, we selected a scale to fit the nonlinear form of the function (Hosmer & Lemeshow, 1989; Selvin, 1996). In this case, we primarily considered quadratic terms. Having identified the appropriate scale for each independent and control variable, we used Kleinbaum's method of backward elimination to decide which variables to include in the final model (Kleinbaum, 1994). According to this method, the analyst begins with a full model and then proceeds to eliminate as many terms as possible, using likelihood ratio tests (analogous to partial *F*-tests in ordinary least-squares regression) to decide which terms contribute significantly to the model's fit. Higher order terms (e.g., interaction terms, squared terms) were eliminated first, then first-order terms. If a higher-order term was kept in the model due to a significant likelihood ratio test, all lower order terms contained in the higher order term were kept in the model as well,

regardless of their *z*-score or likelihood ratio test (Kleinbaum, 1994).

The final step in our modeling strategy was to assess goodness-of-fit. Unlike the case of ordinary least-squares regression, where R^2 provides a widely agreed-upon measure of fit, in logistic regression, there is no consensus measure of goodness-of-fit (Hosmer & Lemeshow, 1989; Kleinbaum, 1994; Pedhazur, 1997; Selvin, 1996). Rather multiple measures are available. For each model, we report two different indices of goodness-of-fit: classification accuracy and Hosmer–Lemeshow's *C* test based on deciles of risk. To calculate classification accuracy, we imposed a threshold on predicted probability scores to generate classifications (e.g., if predicted probability > 0.5, then classify as error). We selected, via systematic search, the threshold that maximized overall accuracy. We reported sensitivity, specificity and overall accuracy at the selected threshold (Hosmer & Lemeshow, 1989).

For the Hosmer–Lemeshow test, we sorted observations into deciles of risk using their predicted probability of error as the sort key. We then compared the observed and expected number of errors and correct responses within each decile of risk (Hosmer & Lemeshow, 1989; Selvin, 1996). Hosmer and Lemeshow's *C* is a χ^2 statistic with 8 degrees of freedom (when deciles are used); the null hypothesis is that the data come from the same distribution (i.e. that the model fits). Plots of predicted versus observed frequencies are provided for each experiment. All statistical tests used $\alpha = 0.05$.

Results and discussion

Each of the 45 participants responded to 160 stimuli, producing 7200 total responses. The mean error rate for blurred, briefly presented typewritten stimuli was 60.6% (4362/7200, std. dev. $= 0.489$). This error rate may seem unrealistically high compared to the rates observed in real world practice settings. Because one subsidiary aim of our investigation was to generate a large number of errors for subsequent analysis, we intentionally inflated the overall error rate (in this experiment and the next) by increasing the amount of degradation and decreasing exposure times. As a result, we now have a large database of actual errors that can be examined in an effort to learn more about the way similarity is represented in the minds of pharmacists. Results of the detailed analysis of errors are presented here only briefly, and a full presentation will be published elsewhere.

On average, each participant incorrectly identified 97 of 160 names (range 51 to 145, median = 95, mode = 99). Table 4 gives the error rate at each level of the independent variables. Table 5 gives the parameter estimates and associated probabilities for the final random effect logistic regression model (i.e., the model arrived at after following Kleinbaum's backward

Table 4
Experiment 1: Error rates at various levels of stimulus frequency, mean neighborhood frequency, and neighborhood density

SF	ND	NF	
		High	Low
High	High	0.47	0.43
	Low	0.41	0.45
Low	High	0.82	0.81
	Low	0.74	0.73
Variable	High	Low	
SF	0.44	0.78	
NF	0.61	0.60	
ND	0.63	0.58	

Table 5
Experiment 1: Parameter estimates for random effects logistic regression model predicting visual perception errors (type-written names)

Variable	Beta	SE	Z
Intercept	0.96	0.13	7.61**
SF	−0.78	0.03	−25.29**
NF	0.06	0.05	1.23
ND	0.10	0.04	2.32*
SF × ND	−0.11	0.03	−3.68**
NF × ND	0.05	0.03	1.69
Trial	−0.01	0.01	−1.91

Note: Terms were kept in the model if likelihood ratio tests of their removal were significant and/or if a higher-level interaction involving a term was significant. In some cases, likelihood ratio tests were significant at $p = 0.05$, but Z-scores (or Wald tests on the parameter estimates) were not. In such cases, likelihood ratio tests are believed to be more reliable, and hence they were used.
*$p < 0.05$, **$p < 0.001$.
−2 log likelihood = 8249.801.

elimination procedure). The model had sensitivity of 72.38%, specificity of 60.29%, and overall accuracy of 67.61% when a predicted probability threshold of 0.55 was used for classification.

Fig. 4 shows the fit between observed and predicted error frequencies at each decile of risk. Although the fit appears to be fairly good, the Hosmer–Lemeshow C test on deciles of risk showed evidence of lack of fit ($X^2(8) = 68.07$, $p = 0.01$). In other words, the hypothesis that predicted and observed distributions came from the same population was rejected. Our decision not to include personal characteristics (e.g., age, gender, ethnicity, visual acuity, years of experience, practice context) as predictors may account for some of the lack
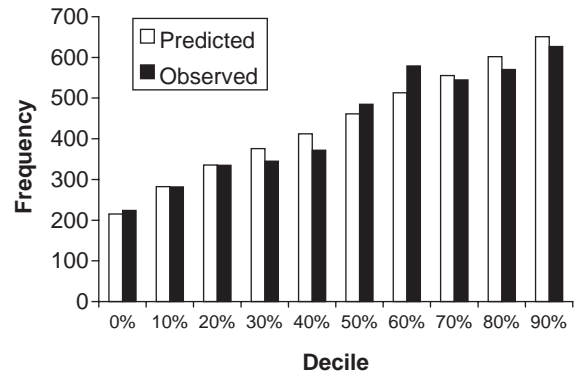
Fig. 4. Fit between predicted and observed number of errors at each decile of risk for Experiment 1.
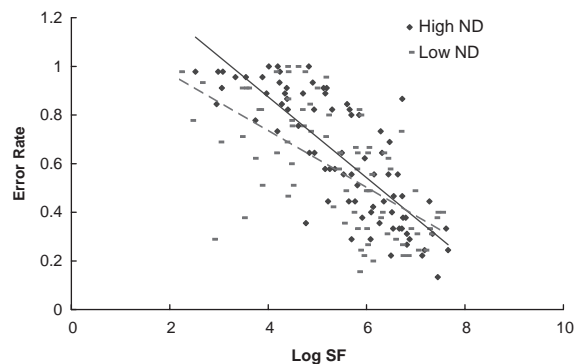

Fig. 5. Interaction between stimulus frequency and neighborhood density in Experiment 1. The solid line represents high neighborhood density. The dotted line represents low neighborhood density.

of fit. Imprecision in the frequency data may also have contributed to lack of fit. In addition, the NAMCS/NHAMCS databases that were used to compute neighborhood characteristics may not have accurately represented each individual pharmacist's personal lexicon. We return to these issues in the Limitations section.

The model revealed a powerful main effect of stimulus frequency. As stimulus frequency increased, the log odds of error sharply decreased (see Fig. 5). The effect of neighborhood density was also significant, although weaker than the stimulus frequency effect. Names in high-density neighborhoods were more difficult to perceive than names in low-density neighborhoods. These effects were contained within a significant two-way interaction between stimulus frequency and neighborhood density (see Fig. 5). The beneficial effect of stimulus frequency on perceptual accuracy was greater for names in high-density neighborhoods than in low-density neighborhoods. Another way of describing the same interaction would be to say that the detrimental

effect of neighborhood density operates primarily on low-frequency names. There was also a significant interaction between neighborhood frequency and neighborhood density. (This term had a non-significant z-score but a significant likelihood ratio test when it was removed from the model. Thus it was retained.) At low levels of neighborhood density, increasing neighborhood frequency reduced the error rate whereas at high levels of neighborhood density, increasing neighborhood frequency increased the error rate (see Fig. 6). Finally, there was a weak but significant effect of trial, with accuracy increasing slightly as participants gained more experience with the task.

Error analysis. Detailed analysis of the errors revealed that there were 234 omission errors and 4128 substitution errors. Of the 4128 substitution errors, 2623 (63.5%) were names of other drugs (e.g., *Indocin*[®] instead of *Indomed*[®], prednisone instead of *Pramasone*[®]). The remaining 1505 (36.5%) substitution errors were spelling errors (e.g., Catapress instead of *Catapres*[®], Cyclogel instead of *Cyclogyl*[®]) or other non-drug responses (e.g., Alena instead of *Alfenta*[®], Seldene instead of *Solatene*[®]). Fig. 7 shows the frequency distribution of substitution errors at different edit distances.

Fig. 7 shows that many substitution errors were within only 1 edit of the target name, suggesting that these may have been typographical errors or misspellings rather than misperceptions. In light of this possibility, we built another model assuming a correct response was any response that fell within 1 edit of the target name. The model was essentially unchanged except that a significant stimulus frequency by neighborhood frequency interaction emerged with the more lenient scoring criterion (details not shown). Given the
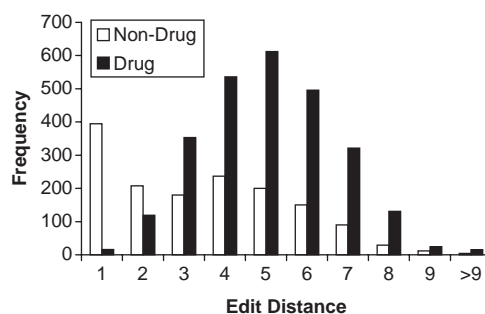
Fig. 7. Histogram of edit distances for both types of substitution errors in Experiment 1. See text for details.

similarity of these models, we will only interpret the initial model.

Results supported Hypothesis 1. All other things being equal, high-frequency drug names (e.g., *Dilantin*[®], *Proventil*[®], *Tagamet*[®]) were perceived more accurately than low-frequency names (e.g., *Hetrazan*[®], *Protaphane*[®], *Antispas*[®]). This is an example of word frequency effect, one of the oldest and most robust effects in all of psycholinguistics (Balota, 1994; Grainger & Dijkstra, 1996; Grainger & Segui, 1990; Monsell, 1991; Whitlow & Cebollero, 1989). The beneficial effects of frequency are presumably due to the higher resting activation of high-frequency names or to decisional biases that favor high-frequency names when stimuli are ambiguous or degraded. Hypothesis 2 was not supported, as there was no reliable main effect of neighborhood frequency. Hypothesis 3 was supported; names in high-density neighborhoods were more likely to be misperceived than names in low-density neighborhoods. For example, all other things being equal, a name like *Betadine*[®], with five neighbors in the NAMCS/NHAMCS database, is more likely to be misperceived than a name like *Cyclogyl*[®], which has only one neighbor.

These main effects, however, must be interpreted in the context of the significant two-way interactions between stimulus frequency and neighborhood density and between neighborhood density and neighborhood frequency. For example, although the main effect of neighborhood density is significant, examination of the interaction with stimulus frequency reveals that density has its effects primarily on low-frequency stimulus names. To understand this effect, recall that the detrimental effects of density are due to competition between similar names; whereas, the beneficial effects of stimulus frequency are due to high levels of resting activation and/or decisional biases. The observed interaction between neighborhood density and stimulus frequency can be explained by noting that the effects of density are of a much smaller magnitude than the effects of stimulus frequency. For high stimulus frequency names, the detrimental effects of density,
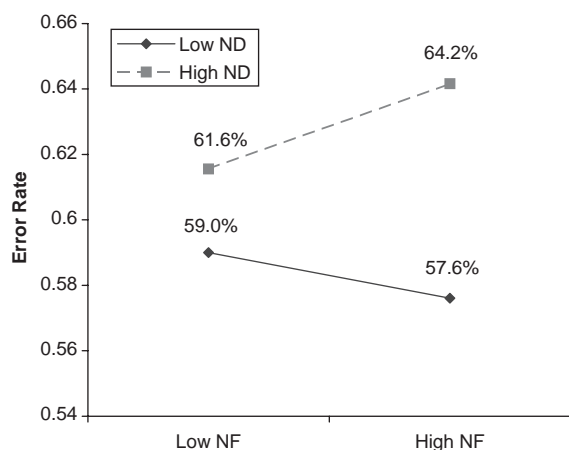

Fig. 6. Interaction between neighborhood frequency and neighborhood density in Experiment 1. The solid line represents low neighborhood density. The dotted line represents high neighborhood density.

although operating, are overwhelmed by the beneficial effects of stimulus frequency. Hence, the observed density effects obtain only at low stimulus frequency. Concretely, competition from the five neighbors of a relatively low-frequency name such as *Moderil*® will make *Moderil*® vulnerable to misperception, but a high-frequency name such as *Nolvadex*® will be relatively unaffected by competition from its five neighbors.

A related explanation can be offered to account for the observed interaction between neighborhood frequency and neighborhood density. In this case, neighborhood frequency increases the error rate in high-density neighborhoods but has the opposite effect in low-density neighborhoods. The more frequent a word is, the more perception is biased in its favor. Thus, one can see neighborhood frequency as amplifying competitive density effects. In contrast, stimulus frequency dampens such effects. If one considers two equally dense neighborhoods, the one with higher frequency neighbors will create more competition with the stimulus word than the one with low-frequency neighbors, since the ability to compete is in proportion to the frequency of a neighbor. This is the effect described by the dotted line in Fig. 6. In contrast, neighborhood frequency has a small, almost negligible, beneficial effect on perceptibility among names in low-density neighborhoods. The small facilitatory effects of neighborhood frequency may be due to increased bottom-up activation coming from neighbor words that share letters with the stimulus word (Andrews, 1997; Zagar & Mathey, 2000). For example, a name like *Altoco*®, which has four neighbors with high average prescribing frequency, will be more vulnerable to competition than a name like *Mantadil*®, which has four relatively low-frequency neighbors.

The next experiment sought to replicate these results in the context of handwritten drug names.

Experiment 2: Perceptual identification of handwritten drug names

Methods

Design

The design of Experiment 2 was identical to that used in Experiment 1 except that in this experiment, handwritten drug names were used as stimuli instead of typewritten names (Brodell et al., 1997). This experiment also used a slightly different type of degradation (described below).

Participants

Participants were 37 licensed, practicing, community pharmacists recruited from the exhibit halls and corridors of the 2000 annual meeting of the National Community Pharmacists Association in San Antonio,

Texas. All participants were adults who consented to participate. All procedures were approved by the local IRB.

Stimulus materials

The stimulus materials were the same 160 names used in Experiment 1 (see Tables 1–3). However, in this experiment, the names were handwritten by 5 practicing physicians from the University of Illinois Hospital and Clinics. The physicians included one fourth-year psychiatry resident, one senior attending physician specializing in occupational medicine, and three internal medicine residents. Physicians were paid US$20 to write each of the 160 names on a multi-page list. Although they were told that the names were to be used in a study of drug name confusion, they were explicitly instructed to write in their normal manner—not to make any extra effort to make their writing neat or legible. The actual handwriting samples used in the experiment were randomly selected from the 5 sets of 160 names. The names were scanned on a Hewlett-Packard ScanJet ADF at 600 dpi and saved as TIFF files. Scanned images were then imported into Adobe Photoshop®, saved in JPEG format, and degraded with Gaussian noise as well as vertical and horizontal graining. The precise procedure for adding noise can be obtained from the authors. (As in Experiment 1, the purpose of added degradation was to ensure that a large number of errors would be generated for subsequent analysis.) Final images were 360 pixels wide and 150 pixels high with 144 pixels per inch resolution. Fig. 8 shows examples of one drug name from each of the five physicians. Due to a programming error, 4 names appeared twice in the experiment (*Mellaril*®, *Pancrease*®, *Restoril*®, and *Zefazone*®) and 4 names never appeared (*Nolvadex*®, *Stelazine*®, *Acular*®, and *Brethancer*®). Data from the second appearance of the repeated names were deleted, and all subsequent analyses were based on 156 names.

Procedure and analysis plan

We used the same procedure and analysis plan as in Experiment 1.

Results and discussion

Thirty-seven participants each responded to 156 names, producing 5772 total responses. The error rate was 45.7% (2637/5772). On average, each participant incorrectly identified 71 of 156 names (std. dev. = 17.76, range, 44 to 131, median = 67). Detailed analysis of errors revealed that there were 202 (7.7%) omission errors and 2435 (92.3) substitution errors. Of the 2435 substitution errors, 1148 (47.1%) involved spelling errors or non-drug names, and 1287 (52.9%) involved other drug names. Fig. 9 shows a histogram of edit distances for drug and non-drug substitution errors.
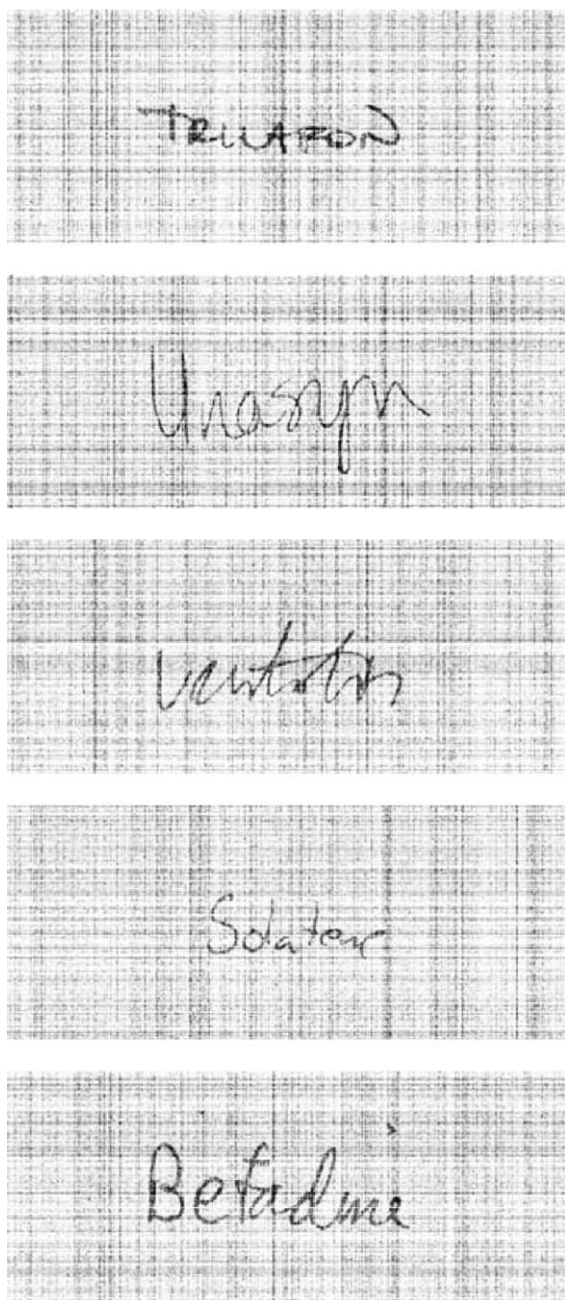
Fig. 8. Examples of degraded handwritten names from five different physicians (from top to bottom: Trilafon®, Unasyn®, Ventolin®, Solatene®, and Betadine®).
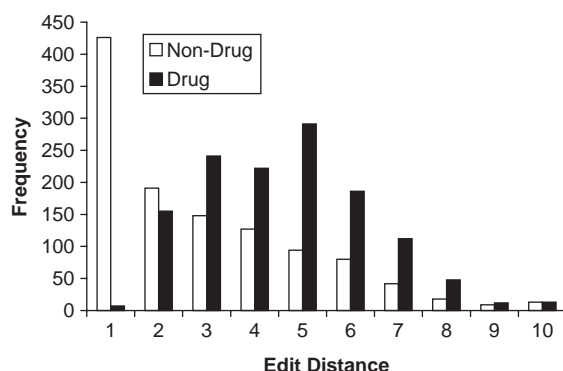

Fig. 9. Histogram of edit distances for both types of substitution errors in Experiment 2. See text for details.

Table 6
Experiment 2: Error rates at various levels of stimulus frequency, mean neighborhood frequency, and neighborhood density

SF	ND	NF	
		High	Low
High	High	0.28	0.36
	Low	0.28	0.31
Low	High	0.68	0.61
	Low	0.59	0.52

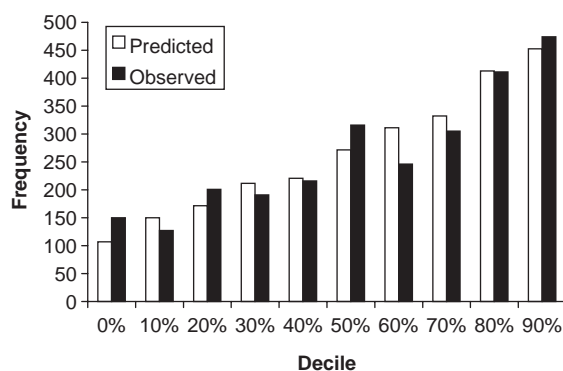| Variable	High	Low
SF	0.31	0.60
NF	0.46	0.45
ND	0.48	0.43

Fig. 10. Fit between predicted and observed number of errors at each decile of risk for Experiment 2.

Table 6 gives the error rates at each level of the independent variables.

The model had sensitivity of 63.37%, specificity of 69.51%, and overall accuracy of 66.70% when a predicted probability threshold of 0.45 was used for classification. Fig. 10 shows the fit between observed and predicted error frequencies at each decile of risk. The Hosmer–Lemeshow C test on deciles of risk led to the rejection of the hypothesis that the observed and predicted distributions came from the same population

(i.e., there was evidence of lack of fit between model and data) ($X^2(8) = 89.79$, $p < 0.01$). The lack of fit was likely due to the same causes identified in Experiment 1.

Table 7

Experiment 2: Parameter estimates for random effects logistic regression model predicting visual perception errors (handwritten names)

Variable	Beta	SE	Z
Intercept	0.13	0.15	0.88
SF	−0.61	0.03	−18.47*
NF	0.10	0.05	1.78
ND	0.19	0.05	3.50*
SF × NF	−0.21	0.04	−4.89*
SF × ND	−0.12	0.03	−3.42*
NF × ND	0.03	0.04	0.73
SF × NF × ND	−0.14	0.02	−6.07*

*$p < 0.001$.

−2 log likelihood = 6949.931.

Table 7 gives the parameter estimates for the random effect logistic regression model. There was a powerful main effect of stimulus frequency, with the log odds of error decreasing sharply as stimulus frequency increased. There was a significant main effect of neighborhood density, with errors being more likely as density increased. The main effect of neighborhood frequency was not significant. There were significant two-way interactions between stimulus frequency and neighborhood density and between stimulus frequency and neighborhood frequency. These were contained within a significant three-way interaction between stimulus frequency, neighborhood frequency, and neighborhood density (see Fig. 11). By definition, a three-way interaction occurs when a two-way interaction differs across levels of a third variable (Keppel, 1991). In this case, stimulus frequency and neighborhood density interacted at high neighborhood frequency but not at low neighborhood frequency. At high neighborhood frequency, stimulus frequency had a stronger dampening effect on names from high as opposed to low-density
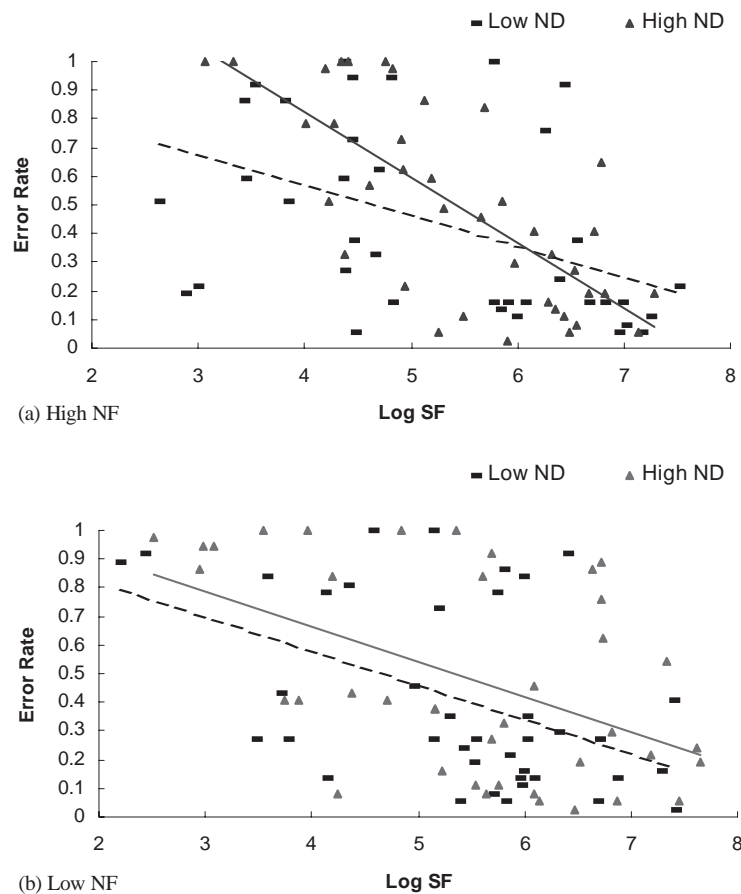

(a) High NF

(b) Low NF

Fig. 11. Three-way interaction between SF, NF, and ND in Experiment 2. Solid lines represent high ND. Dotted lines represent low ND. Panel (a) illustrates a two-way interaction between SF and ND at high NF, but this interaction is not present at low NF (panel b).

neighborhoods. At low neighborhood frequency, the effect of stimulus frequency was the same regardless of neighborhood density. In other words, high prescribing frequency of a target name is most protective against error in high-frequency, high-density neighborhoods, where competition is fiercest. This makes sense given that the ability to compete in perception is mostly a function of frequency.

Results supported Hypothesis 1. All things being equal, common names (e.g., *Ativan*®) were more accurately perceived than rare names (e.g., *Protophane*®). Hypothesis 2 was not supported. The main effect of neighborhood frequency was not reliably greater than zero. Hypothesis 3 was supported. Neighborhood density significantly increased the likelihood of error, i.e., the more neighbors a name has, the harder that name is to correctly identify. One must be cautious in interpreting the main effects, however, since they were contained within significant two- and three-way interactions. The three-way interaction is still in need of explanation.

At high neighborhood frequency, the interaction between stimulus frequency and neighborhood density is the same as that which was observed for typewritten names, namely, that the effect of stimulus frequency was greater on names from high as opposed to low-density neighborhoods. To put it another way, in high-frequency neighborhoods, high neighborhood density had detrimental effects on low-frequency names but beneficial effects on high-frequency names. In low-frequency neighborhoods, there was no such interaction. Instead, low-frequency names showed only the main effects of neighborhood density and stimulus frequency. Again we interpret this as neighborhood frequency amplifying the effects of density. In low-frequency neighborhoods, there is a density effect, but because the neighbors are relatively rare names, they exert relatively little competitive inhibition on the target name—hence, the small main effect of neighborhood density in low-frequency neighborhoods. In contrast, neighborhoods that are dense with high-frequency names compete very strongly with the target name, creating the large density effect, especially on rarely prescribed target names.

General discussion

Drug name confusions are a persistent source of medication errors and an ongoing threat to patient safety. In spite of all that has been written and all the remedial steps that have been taken recently, new pairs of confusing names continue to appear regularly (Institute for Safe Medication Practices, 2002). One way to minimize the incidence of these errors is to equip decision-makers in the pharmaceutical industry and the FDA with tools to make better name-approval decisions. Such tools should be empirically validated and based on principles of human factors engineering. When it comes to drug name confusion, one of the fundamental human factors is visual perception (others are short-term memory and auditory perception).

The present work provides new evidence regarding what factors contribute to visual misidentifications of drug names. Consistent with current theory in cognitive psychology, the present results indicate that highly frequent drug names are less likely to be misperceived than less frequent drug names. Also consistent with current theory, drug names that are similar to many other drug names are more likely to be misperceived than drug names that are similar to relatively few other drug names. These findings have important implications both for current pharmaceutical practice and for the regulation of proposed new drug names. The implications are not obvious, however, and in order to understand them, one must consider the ultimate goal of those who might use the types of predictive models described above.

The ultimate goal of decision-makers ought to be to minimize harm (i.e., to maximize patient safety). Doing so involves more than just predicting the probability that a given name will be misperceived, which is all we have done here. Harm is a function of the probability of error, the number of opportunities for error and the severity of the consequences of an error:

$$\text{Harm} = \text{Probability of Error} \\ \times \text{Number of Opportunities for Error} \\ \times \text{Severity of Each Error}.$$

Although our work clearly documents that rare names are more likely to be misidentified than common names, it does not necessarily follow that rare names pose the most risk of harm. For example, imagine Drug A is prescribed 1 million times per year and has a predicted probability of error of 0.05. Imagine that Drug B is prescribed 10,000 times per year and has a predicted probability of error of 0.5. Alone, the probabilities of error suggest that Drug B is 10 times more likely to be misidentified than Drug A. One might be tempted to conclude that Drug B is therefore 10 times more hazardous, but this would be a misguided conclusion. Given these hypothetical prescribing frequencies and error probabilities, one would predict 0.05×1 million $= 50{,}000$ errors for Drug A and $0.5 \times 10{,}000 = 5000$ errors for Drug B. The arithmetic is simple, but the implications are somewhat counter-intuitive. The drug with the lower probability of error (the frequently prescribed Drug A) is likely to be involved in many more errors than the much harder to perceive Drug B, by virtue of the much greater number of opportunities for error involving Drug A. Even this is not the whole

picture, because the severity of the consequences of an error has not been factored in. Five thousand fatal errors would clearly cause more harm than 50,000 mild allergic reactions.

The upshot of all this is that to make the drug name lexicon safer, one must focus on harm reduction, and that harm reduction involves more than just the probability of error. It also involves the number of opportunities for error and the amount of damage caused by each error. It is worth noting that, in most cases, harm is defined in terms of injury, disability, and death to patients who get the wrong drug. But sometimes harm may be defined in terms of commercial damages, lost sales, trademark infringement, or trademark dilution. Regulatory agencies like the US Food and Drug Administration are likely to use the former definition (injury) while drug companies, the courts and other regulators like the US Patent and Trademark Office may use the latter definition (commercial harm). In neither case are valid and reliable measures for quantifying the magnitude of harm very well developed. This is clearly an area where more work needs to be done.

Comparison of results for handwritten and typewritten names

The main effects of stimulus frequency and neighborhood density in Experiment 2 were in the same direction and of comparable magnitude as what was observed in Experiment 1. Thus, the effects of stimulus frequency and neighborhood density appear, for the most part, to apply to both handwritten and typewritten names. Similarly, the main effect of neighborhood frequency is null for both handwritten and typewritten names. The results obtained with the handwritten materials serve two important purposes. First, they serve as a partial replication of the results obtained with typewritten materials, thus establishing the robustness of those results. Second, they extend those results to the more realistic domain of handwritten stimuli of the kind that are more likely to be encountered in everyday practice.

Limitations

Only pharmacists participated in these experiments. To strengthen generalizations about other populations, experiments should be repeated using physicians, nurses, and lay people as participants. The names used as stimuli were mostly 3-syllable brand names. Generalizing from these results to generic names or names of different lengths may or may not be warranted. The type of stimulus degradation used, which combined degraded images with short presentation times, may not perfectly simulate the real-world perceptual situation, where names may be more or less degraded, handwriting may be better or worse, and pharmacists may spend more or less time examining ambiguous names. The task used here may also be somewhat artificial in that we encouraged participants to generate a guess whenever possible whereas in the real world, a pharmacist would presumably seek verification of an ambiguous name rather than guessing. Furthermore, the task took place in a quiet room, with none of the distractions that would be present in the real world. No attempt was made to model variation that may have been due to a participant's age, years of experience, practice context, clinical specialty, or level of visual acuity, nor were differences in the legibility of the handwritten names taken into account. The absence of these variables may, in part, explain the lack of fit between observed and predicted data. In addition, the lexicon used here included names only from the NAMCS and NHAMCS databases. Limitations in the precision of NAMCS/NHAMCS frequency data may have affected the precision of our frequency-based parameter estimates. Use of a different reference database would alter the neighborhood characteristics of the stimulus names and the resulting neighborhood effects. Finally, the observed error rates primarily reflect characteristics of the experimental task such as the brief exposure duration and the nature and degree of stimulus degradation. The absolute error rates observed here (45–60%) are perhaps an order of magnitude higher than what one might expect in a real-world practice setting. Hence, interpretation of these results should focus on the trends in error rates as a function of frequency and density, not on the absolute error rates. To put it more formally, interpretations should focus on the slopes of the models, not their y-intercepts.

Conclusion

Pharmacists' visual perception of briefly presented, blurry drug names is affected in predictable ways by objectively measurable properties of names (e.g., prescribing frequency, neighborhood frequency, and neighborhood density). All other things being equal, rare names and names in high-density neighborhoods are more likely to be misperceived than common names and names from low-density neighborhoods. The main effects of neighborhood frequency are negligible, although this factor does interact in important ways with neighborhood density and stimulus frequency (i.e., neighborhood frequency tends to amplify the effects of density). The complex interactions observed for both handwritten and typewritten stimuli suggest that all three factors must be simultaneously taken into account when assessing visually intelligible or visually confusing drug names. Models such as the ones presented here, or refinements thereof, can be used for this purpose. People

responsible for coining and approving drug names can and should consider the factors identified here when they make decisions. However, the models presented here only pertain to the probability of error. In order to minimize harm, which ought to be the goal of practitioners, industry, and government alike, one must also take into account the number of opportunities for error and the severity of the consequences of each error. We hope to have facilitated the process of harm reduction by offering insight into the factors that affect the probability of error. Traditional pharmaceutical marketing techniques should be able to estimate the number of opportunities for error (i.e., the prescribing frequency of a drug), and further work is needed to quantify the harm associated with each error.

References

Andrews, S. (1997). The effect of orthographic similarity on lexical retrieval: Resolving neighborhood conflicts. *Psychonomic Bulletin & Review*, *4*(4), 439–461.

Balota, D. A. (1994). Visual word recognition: The journey from features to meaning. In M. A. Gernsbacher (Ed.), *Handbook of psycholinguistics* (pp. 303–358). San Diego, CA: Academic Press.

Brodell, R. T., Helms, S. E., KrishnaRao, I., & Bredle, D. L. (1997). Prescription errors. Legibility and drug name confusion. *Archives of Family Medicine*, *6*, 296–298.

Carreiras, M., Perea, M., & Grainger, J. (1997). Effects of orthographic neighborhood in visual word recognition: Cross-task comparisons. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *23*(4), 857–871.

Cohen, J. D., MacWhinney, B., Flatt, M., & Provost, J. (1993). PsyScope: A new graphic interactive environment for designing psychology experiments. *Behavioral Research Methods, Instruments, and Computers*, *25*(2), 257–271.

Coltheart, M., Davelaar, E., Jonasson, J. T., & Besner, D. (1977). Access to the internal lexicon. In S. Dornic (Ed.), *Attention and performance VI* (pp. 535–555). New York, NY: Academic Press.

Dell, G. S., Schwartz, M. F., Martin, N., Saffran, E. M., & Gagnon, D. A. (1997). Lexical access in normal and aphasic speakers. *Psychological Review*, *104*, 801–838.

Grainger, J., & Dijkstra, T. (1996). Visual word recognition: Models and experiments. In T. Dijkstra, & K. de Smedt (Eds.), *Computational psycholinguistics* (pp. 139–165). Bristol, PA: Taylor & Francis.

Grainger, J., & Jacobs, A. M. (1996). Orthographic processing in visual word recognition: A multiple read-out model. *Psychological Review*, *103*(3), 518–565.

Grainger, J., & Segui, J. (1990). Neighborhood frequency effects in visual word recognition: A comparison of lexical decision and masked identification latencies. *Perception & Psychophysics*, *47*, 191–198.

Havens, L. L., & Foote, W. E. (1963). The effect of competition on visual duration threshold and its independence of stimulus frequency. *Journal of Experimental Psychology*, *65*, 6–11.

Hedeker, D. (1999). *MIXOR/MIXREG Website* [HTML page]. University of Illinois at Chicago. Available: http://www.uic.edu/~hedeker/mix.html [1999, June 4].

Hedeker, D., & Gibbons, R. D. (1994). A random-effects ordinal regression model for multilevel analysis. *Biometrics*, *50*(4), 933–944.

Hedeker, D., & Gibbons, R. D. (1996). MIXOR: A computer program for mixed-effects ordinal regression analysis. *Computer Methods and Programs in Biomedicine*, *49*, 157–176.

Hosmer, D. W., & Lemeshow, S. (1989). *Applied logistic regression*. New York: Wiley.

Howes, D. H., & Solomon, R. L. (1951). Visual duration threshold as a function of word probability. *Journal of Experimental Psychology*, *41*, 401–410.

Institute for Safe Medication Practices. (2002). *ISMP medication safety alert!* (7(7)). Huntingdon Valley, PA: Institute for Safe Medication Practices.

Jacobs, A. M., & Grainger, J. (1994). Models of visual word recognition—Sampling the state of the art. *Journal of Experimental Psychology: Human Perception and Performance*, *20*(6), 1311–1334.

Keppel, G. (1991). *Design and analysis: A researcher's handbook*. Englewood Cliffs, NJ: Prentice-Hall.

Kleinbaum, D. G. (1994). *Logistic regression: A self-learning text*. New York: Springer.

Kohn, L. T., Corrigan, J., & Donaldson, M. S. (Eds.). (2000). *To err is human: Building a safer health system*. Washington, DC: Institute of Medicine.

Lambert, B. L., Chang, K. Y., & Lin, S. J. (2001a). Descriptive analysis of the drug name lexicon. *Drug Information Journal*, *35*, 163–172.

Lambert, B. L., Chang, K. Y., & Lin, S. J. (2001b). Effect of orthographic and phonological similarity on false recognition of drug names. *Social Science & Medicine*, *52*, 1843–1857.

Lambert, B. L., Chang, K.-Y., Lin, S.-J., & Gupta, P. (2000). Effect of prescribing frequency, neighborhood frequency, and neighborhood density on pharmacists' visual perception of drug names. *PharmSci*, *2*, S–1276.

Lambert, B. L., Lin, S.-J., Gandhi, S. K., & Chang, K.-Y. (1999). Similarity as a risk factor in drug name confusion errors: The look-alike (orthographic) and sound-alike (phonological) model. *Medical Care*, *37*(12), 1214–1225.

Luce, P. A., & Pisoni, D. B. (1998). Recognizing spoken words: The neighborhood activation model. *Ear & Hearing*, *19*(1), 1–36.

Luce, P. A., Pisoni, D. B., & Goldinger, S. D. (1990). Similarity neighborhoods of spoken words. In G. T. M. Altmann (Ed.), *Cognitive models of speech processing: Psycholinguistic and computational perspectives* (pp. 122–147). Cambridge, MA: MIT Press.

Luce, R. D. (1959). *Individual choice behavior*. New York: Wiley.

Mathey, S. (2001). L'influence du voisinage orthographique lors de la reconnaissance des mots ecrits [The effect of orthographic neighbourhoods on visual recognition of written words]. *Revue Canadienne de Psychologie Experimentale*, *55*(1), 1–23 (in French).

McClelland, J. L., & Rumelhart, D. E. (1981.). An interactive activation model of context effects in letter perception: Part I. An account of basic findings. *Journal of Experimental Psychology: General*, *114*(159-188).

Monsell, S. (1991). The nature and locus of word frequency effects in reading. In D. Besner, & G. W. Humphreys (Eds.), *Basic processes in reading: Visual word recognition* (pp. 148–197). Hillsdale, NJ: Erlbaum.

National Center for Health Statistics. (2001a). *Ambulatory health care data: NAMCS description* [Web page]. National Center for Health Statistics. Available: http://www.cdc.gov/nchs/about/major/ahcd/namcsdes.htm [2001, May 31, 2001].

National Center for Health Statistics. (2001b). *Ambulatory health care data: NHAMCS description* [Web page]. National Center for Health Statistics. Available: http://www.cdc.gov/nchs/about/major/ahcd/nhamcsds.htm [2001, May 31, 2001].

Newbigging, P. L. (1961). The perceptual reintegration of frequent and infrequent words. *Canadian Journal of Psychology*, *15*, 123–132.

Pedhazur, E. J. (1997). *Multiple regression in behavioral research: Explanation and prediction*. New York: Harcourt Brace.

Perea, M., & Rosa, E. (2000). The effects of orthographic neighborhood in reading and laboratory word identification tasks. *Psicologica*, *21*, 327–340.

Rumelhart, D. E., & McClelland, J. L. (1982). An interactive activation model of context effects in letter perception: Part II. The contextual enhancement effect and some tests and extensions of the model. *Psychological Review*, *89*, 60–94.

Selvin, S. (1996). *Statistical analysis of epidemiological data*. New York: Oxford University Press.

Snodgrass, J. G., & Mintzer, M. (1993). Neighborhood effects in visual word recognition: Facilitatory or inhibitory? *Memory and Cognition*, *21*(2), 247–266.

Solomon, R. L., & Postman, L. (1952). Frequency of usage as a determinant of recognition threshold for words. *Journal of Experimental Psychology*, *43*, 195–210.

Stephen, G. A. (1994). *String searching algorithms*. River Edge, NJ: World Scientific.

US Pharmacopeia. (1995). *USP quality review* (49). Rockville, MD: US Pharmacopeia.

US Pharmacopeia. (1996). *USP quality review*. Rockville, MD: US Pharmacopeia.

US Pharmacopeia. (2001). *USP quality review* (76). US Pharmacopeia. Available: http://www.usp.org/reporting/review/qr76.pdfJuly 19, 2001], March.

US Food and Drug Administration. (2001). *Medication errors*, [Web Page]. Available: http://www.fda.gov/cder/handbook/mederror.htm [2001, April 23, 2001].

Whitlow, J. W., & Cebollero, A. (1989). The nature of word frequency effects in perceptual identification. *Journal of Experimental Psychology: Learning, Memory, and Cognition*, *15*(4), 643–656.

Zagar, D., & Mathey, S. (2000). When WORDS with higher-frequency neighbours become words with no higher-frequency neighbour (or how to undress the neighborhood frequency effect). In A. Kennedy, R. Radach, D. Heller, & J. Pynte (Eds.), *Reading as a perceptual process* (pp. 23–46). New York, NY: Elsevier Science.